

A NOTE ON THE LIMITING DISTRIBUTION OF "JACKSON'S ESTIMATE" OF  
AVERAGE MORTALITY RATE

EU-130-M

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ABSTRACT

If  $n_x$  is the number of  $x$ 's in a sample of size  $n$  from the distribution

$$f(x) = s_1 s_2 \cdots s_x / (1 + s_1 + s_1 s_2 + \cdots + s_1 s_2 \cdots s_k) \quad x=0, 1, \cdots, k$$

then the maximum likelihood estimator of  $s_i$  is  $S_i = n_i / n_{i-1}$ . The asymptotic distribution of the  $S$  vector is the multivariate normal with mean  $s$  and a covariance matrix estimated by

$$\hat{\sigma}_{ii} = \frac{n_i^2}{n_{i-1}^2 (n_0 + n_1 + \cdots + n_{i-1})} \left[ \frac{n_0 + \cdots + n_i}{n_i} + \frac{n_0 + \cdots + n_{i-1}}{n_{i-1}} \right]$$

$$\hat{\sigma}_{i,i+1} = - \frac{n_{i+1}}{n_{i-1} n_i}$$

All other covariances are zero in this asymptotic distribution.

The variance of the asymptotic normal distribution of the estimate of the average  $\bar{S} = (S_1 + \cdots + S_k) / k$  is readily obtained from these results.

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Let  $X$  be a chance variable having the probability distribution

$$f(x) = s_1 s_2 \cdots s_x / (1 + s_1 + s_1 s_2 + \cdots + s_1 s_2 \cdots s_k)$$

for  $x=0, 1, \dots, k$  and suppose  $X_1, \dots, X_n$  are independent and distributed as  $X$ .

Let

$$n_x = \text{number } X_i \text{ which are equal to } x$$

then the log likelihood of  $X_1, \dots, X_n$  may be written

$$\begin{aligned} \log L = & (n_1 + \cdots + n_k) \log s_1 + (n_2 + \cdots + n_k) \log s_2 + \cdots + n_k \log s_k \\ & - n \log(1 + s_1 + s_1 s_2 + \cdots + s_1 s_2 \cdots s_k) \end{aligned}$$

The maximum likelihood equations then become

$$\frac{\partial \log L}{\partial s_i} = \frac{n_1 + \cdots + n_k}{s_i} - \frac{n(s_1 \cdots s_i + s_1 \cdots s_{i+1} + \cdots + s_1 \cdots s_k)}{s_i(1 + s_1 + s_1 s_2 + \cdots + s_1 s_2 \cdots s_k)} = 0$$

or, letting

$$\alpha_i = (s_1 \cdots s_i + s_1 \cdots s_{i+1} + \cdots + s_1 \cdots s_k) / (1 + s_1 + s_1 s_2 + \cdots + s_1 s_2 \cdots s_k)$$

then

$$\frac{\partial \log L}{\partial s_i} = \frac{n_1 + \cdots + n_k}{s_i} - \frac{n\alpha_i}{s_i} = 0$$

Consequently, the maximum likelihood estimator of  $\alpha_i$  is  $A_i = (n_1 + \cdots + n_k)/n$  and since

$$s_i = \frac{\alpha_i - \alpha_{i+1}}{\alpha_{i-1} - \alpha_i}$$

then the maximum likelihood estimator of  $s_i$  is

$$S_i = \frac{A_i - A_{i+1}}{A_{i-1} - A_i} = \frac{n_i}{n_{i-1}}$$

as was mentioned by Chapman and Robson [1].

The average of the  $k$  estimators  $S_1, \dots, S_k$  is referred to as Jackson's estimate [2] and is widely used in fisheries work as an estimator of average annual mortality rate. We shall now compute the covariance matrix of the limiting distribution of  $S_1, \dots, S_k$  and from this obtain the variance of the limiting distribution of

$$\bar{S} = \frac{1}{k}(S_1 + \dots + S_k)$$

The inverse elements of the asymptotic covariance matrix are

$$\sigma^{ii} = -E \frac{\partial \log L^2}{\partial s_i^2} = E \left( \frac{n_i + \dots + n_k}{s_i^2} \right) - \frac{n\alpha_i^2}{s_i^2} = \frac{n\alpha_i(1-\alpha_i)}{s_i^2}$$

and, for  $i < j$ ,

$$\sigma^{ij} = -E \frac{\partial \log L^2}{\partial s_i \partial s_j} = \frac{n\alpha_j(1-\alpha_i)}{s_i s_j}$$

This matrix  $\|\sigma^{ij}\|$  is of the same form considered in an earlier note [3], and applying those results directly to this problem we obtain the inverted matrix

$$\sigma_{ii} = \frac{s_i^2}{n(1-\alpha_i)} \left[ \frac{1-\alpha_{i+1}}{\alpha_i - \alpha_{i+1}} + \frac{1-\alpha_{i-1}}{\alpha_{i-1} - \alpha_i} \right] \quad \text{for } i=1, \dots, k$$

$$\sigma_{i,i+1} = - \frac{s_i s_{i+1}}{n(\alpha_i - \alpha_{i+1})} \quad \text{for } i=1, \dots, k-1$$

and all other covariances are zero. The variance of Jackson's estimate is therefore asymptotically equal to

$$\text{var}(\bar{S}) = \frac{1}{k^2} \left[ \sum_{i=1}^k \sigma_{ii} + 2 \sum_{i=1}^{k-1} \sigma_{i,i+1} \right]$$

Maximum likelihood estimators of these variances and covariances are given by

$$\hat{\sigma}_{ii} = \frac{n_i^2}{n_{i-1}^2 (n_0 + n_1 + \dots + n_{i-1})} \left[ \frac{n_0 + \dots + n_i}{n_i} + \frac{n_0 + \dots + n_{i-1}}{n_{i-1}} \right]$$

$$\hat{\sigma}_{i,i+1} = - \frac{n_{i+1}}{n_{i-1} n_i}$$

References

- [1 ] Chapman, D. G. and D. S. Robson. The analysis of a catch curve. Biometrics Vol. 16:354-68, 1960.
  
- [2 ] Jackson, C. H. N. The analysis of an animal population. J. Animal Ecology 8:238-46, 1939.
  
- [3 ] Robson, D. S. Maximum likelihood estimation of a sequence of annual survival rates from a capture-recapture series. Biometrics Unit Mimeograph Series BU-116-M, Cornell University, Ithaca, New York. 1960.